

The Private Return on Education and How to Solve the Endogeneity Problem: Case Indonesia

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Abstract—This paper discusses the return on education using the Mincer model. The Mincer equation is a log natural income associated with years of school completion and work experience. One of the problems arising from the Mincer equation using the OLS (Ordinary Least Square) method is biased estimation result. The magnitude of the bias collected by previous researcher's ranges from 0.7% -9.4%. Bias can occur due to sample selection and endogeneity problems. Problem of sample selection will be overcome with mills ratio or Heckit method while the endogeneity problem in this paper will be overcome by IV (Instrumental Variable) method. Instrumental variables must have a provision: the variables used as the instrument have no relation with the dependent variable, and the variable has correlated with the independent variable that is considered endogenous. Variable used as instrument is parent education. However, in this paper, I use pooling data from 1993-2014 with Indonesia Family Live Survey. Empirical results show the difference between return on education value using OLS and IV methods. Results with IV give a larger return value than OLS. The return value bias for the Indonesian case is still in tolerance.

Keywords—return on education; mincer; pooling data; sample selection; instrumental variable

I. INTRODUCTION

Return to education is an economic advantage of one's investment in education. There are three different sides in defining the return of education, namely: private return, social return, and labor productivity return [1]. The famous return to education model is the mincer model [2]. This paper focuses on private return and uses the Mincer equation to calculate the estimated return on education.

The Mincer equation allows for bias due to selection, unobserved ability, and endogeneity. So the education coefficient can't be interpreted as a pure measure of educational impact. The coefficient of educational estimate becomes $E(\hat{\beta}_1) \neq \beta_1$ [3]. Some ways to overcome the bias include adding proxy variables to the equations, estimating models using twin data, using two-stage estimates, and using instrument variables [4].

A study using twins' data provides estimation results that minimize bias, while when using IV (Instrumental Variable)

method the estimation results obtained are greater than the estimation results using the OLS (Ordinary Least Square) method. In this study the instruments used were family backgrounds or the school system [5].

Until now there hasn't been research that can conclude the magnitude of the bias that occurs due to using the OLS method in the Mincer equation [6]. Nevertheless, several studies have collected the results of previous studies on estimation of return on education using m IV method. The summary collected by the researchers provides an overview of the magnitude of the bias ranging from 0.7% - 9.4% [5,7].

In general, the estimation results carried out by several studies using IV method have an estimation value that is greater than the OLS method [8-13].

Studies on the return on education case of Indonesia have been carried out by several researchers. One researcher used a two-step regression method with a multinomial logit used in the first step, then the estimation results were included in the main model. This study uses IFLS1 to IFLS3 [14]. Other study used OLS and Heckit methods. It used SUSENAS 2005 data [15]. Other study used OLS, Heckit and fixed effect methods. Its study results that estimation of return on education with the Heckit and the Fixed Effect methods are smaller than the OLS method [16]. Other study used OLS as its basic methodology and compares it with two estimation steps from Heckman (Heckit's method) [17].

The benefits of this research are to find solutions to problem solving of the return on education estimation of the Mincer model for the case in Indonesia. In addition, to determine the amount of bias that occurs due to the use of OLS. Data for twins in Indonesia have not been collected well, so this paper uses the IV method to overcome the bias due to unobserved ability and endogeneity. The Heckit method will be used to see if there is a bias selection in this case.

However, the difference in the settlement of this case with previous researchers is the use of IV methods and the use of pooling cross-section data. It is hoped that this paper will provide additional information on the large bias that occurs and also shows the impact of time on changes in the value of return on education.

II. METHOD

A. Framework Mincer Equation

The simple model of school decision on partial balance describes the tradeoff in human capital investment. The framework of the income and education equation model refers to Acemoglu and Autor [18]. Every human being will try to maximize the utility function $u(c)$ along the planning horizon T ($T = \infty$) with the positive discount rate ($\rho > 0$), and the rate of constant death rate ($v \geq 0$), which is reflected in the equation:

$$\max \int_0^T \exp(-(\rho + v)t) u(c(t)) dt.$$

Individuals have the human capital growth function $\dot{h}(t) = G(t, h(t), s(t))$, where each individual is within the interval S , $s(t) = 1$ at school, and $s(t) = 0$ when finished school. At the end of the school interval, the individual will have school level $h(S) = \eta(S)$, where $\eta(\cdot)$ is an increasing, continuous and concave function.

When $t \in [S, \infty)$, the accumulation of human capital throughout life (due to individual work) has the equation $\dot{h}(t) = g_h h(t)$, with the growth of human capital ($g_h \geq 0$). Individuals have an exponential wage growth with the equation $\dot{w}(t) = g_w w(t)$, with growth wages (g_w) and wage conditions at the beginning $w(0) > 0$.

Suppose that $g_w + g_h < r + v$, optimal school decision is a maximal solution of $\max_S \int_S^\infty \exp(-(\rho + v)t) w(t) h(t) dt$. The equation is equivalent to

$$\max_S \frac{\eta(S) w(0) \exp(-(\rho + v - g_w)S)}{r + v - g_h - g_w} \tag{2}$$

First order condition equation (2) to S is

$$\frac{\eta'(S^*)}{\eta(S^*)} = r + v - g_w \tag{3}$$

Equation (3) shows a high value at r (interest rate) and v (short time horizon) will decrease human capital investment. While the high growth value of wages increases the value of human capital and will encourage even greater investment.

By integrating both sides of equation (3) with respect to the variable S we obtain.

$$\ln \eta(S^*) = constant + (r + v - g_w)S^* \tag{4}$$

Earnings of a worker in the age of $\tau \geq S^*$ in the labor market at time t will be obtained.

$$W(S, t) = \exp(g_w t) \exp(g_h(t - S)) \eta(S). \tag{5}$$

By providing the logarithm of equation (5) and substituting in equation (4) the income of the worker will be in the form.

$$\ln W(S^*, t) = constant + (r + v - g_w)S^* + g_w t + g_h(t - S^*). \tag{6}$$

Where $(t - S^*)$ is the worker's experience (time after school). If we use cross sectional to compare between workers, then time trend $g_w t$ will be constant, so we get a canonical Mincer equation in which log wage is proportional to school and experience. Equation (6) can be simplified to be.

$$\ln W_i = constant + \beta_s S_i + \beta_e E_i, \tag{7}$$

Where i is an individual i . The coefficient value of β_s will be positive if $r + v > g_w$. It will produce the coefficient $\beta_e < \beta_s$. Experience variable (E) is individual potential experiences derived from calculations (age - years of education - 7).

B. Empirical Strategy

The Mincer model used is shaped

$$\ln w_i = \beta_0 + \beta_1 S_i + \beta_2 E_i + \beta_3 E_i^2 + \gamma X + \varepsilon, \tag{8}$$

where X is a variable vector of individual characteristics.

This study uses pooling cross section. The advantages of using pooling cross-sections method are increasing sample size, and resulting in more precise estimates. Other benefits can identify the impact of time, so as to see whether the relationship between education and income changes over time [3]. The model pooling cross section is in the form of:

$$\ln w_i = \beta_0 + \beta_1 S_i + \sum_{k=1}^4 \rho_k S_i T_k + \beta_2 E_i + \beta_3 E_i^2 + \sum_{k=1}^4 T_k + X\gamma + \varepsilon, \tag{9}$$

The sample used in this research is the income workers. The selection of samples like this of course tends to the problem of sample selection. The problem of sample selection can be overcome by a two stage method introduced by Heckman [19]. The procedure for completion of sample selection correction as follows: (1) using all observations to find the probability of working by using probit model, then count the inverse Mills ratio for each observed sample. (2) Use the selected sample by entering the mills ratio value into the main model. Significant mills ratio coefficients indicate a selection problem bias, so the estimation of the coefficient of the main model with OLS will be biased [3]. This two stage method is known as the Heckit method.

Instrumental variable methods are used to overcome possible bias due to endogeneity. The problem of the instrument variable method is to find the right instrument for the education variable that is considered endogenous. Some studies used family backgrounds as instruments, such as: mother's education [5], father's education or the education of

her siblings [20]. Other instruments used for educational variables include "birth quarter interacted with year of birth" [8,9], "closest distance to campus" [10], "age cohort indicator" [21], "dummy reform of the school system at the age of 13 years" [22]. In this paper, I use parent education (father or mother) as an instrument of the educational variables.

III. RESULTS

This paper uses IFLS data from first wave (1993) to fifth wave (2014). The respondents selected who have ages between 15 and 65 years. The number of observations in this data is shown in Table 1. All observations after deduction of education missing and ages limits totaled 114,618 and those with income data were 68,060. Decrease in observation due to missing data on one's income.

The statistical summary of the data obtained is shown in Table 2. The average education in Indonesia increases from year to year. The average education in 1993 was 5,365 years and in 2014 it was 9,043 years.

The probability model working or not working from all respondents needs to be made first. The descriptive analysis of the probability model is shown in table 3. This probability data is used to estimate the Probit model to overcome the selection bias.

The average probability of work increases every year except in 1997. The average probability of employment in 2014 increased by around 17.2% compared to 1993. While in 1997 it decreased by around 9.6% compared to 1993.

TABLE I. SUMMARY OF OBSERVATIONS

	Observation	
	<i>Dumped</i>	<i>Remainder</i>
<i>IFLS1</i>		
Individuals who answered Book III		14,418
Observation discarded due		
Education <i>missing</i>	21	14,397
Age > 65	1,248	13,149
Income <i>missing</i> and <i>outlier</i>	5,717	7,432
<i>IFLS2</i>		
Individuals who answered Book III		19,910
Observation discarded due		
Education <i>missing</i>	45	19,865
Age > 65	1,264	18,601
Income <i>missing</i> and <i>outlier</i>	9,096	9,505
<i>IFLS3</i>		
Individuals who answered Book III		25,490
Observation discarded due		
Education <i>missing</i>	26	25,464
Age > 65	1,666	23,798
Income <i>missing</i> and <i>outlier</i>	9,954	13,844
<i>IFLS4</i>		
Individuals who answered Book III		29,059
Observation discarded due		
Education <i>missing</i>	3	29,056
Age > 65	1,734	27,322
Income <i>missing</i> and <i>outlier</i>	11,186	16,135
<i>IFLS5</i>		
Individuals who answered Book III		34,464
Observation discarded due		
Education <i>missing</i>	33	34,431
Age > 65	2,683	31,748
Income <i>missing</i> and <i>outlier</i>	10,604	21,144

Source: IFLS1, IFLS2, IFLS3, IFLS4, and IFLS5

TABLE II. SUMMARY OF WORK PROBABILITY MODEL

Data	Variable	Observation	Average
<i>IFLS1</i>	Worked Probability	13,149	0.565
	Education		5.365
	Age		40.074
	Men		0.453
<i>IFLS2</i>	Worked Probability	18,601	0.511
	Education		6.755
	Age		34.870
	Men		0.456
<i>IFLS3</i>	Worked Probability	23,798	0.582
	Education		7.484

Table 2. Cont.

	Age		33.733
	Men		0.474
<i>IFLS4</i>	Worked Probability	27,322	0.590
	Education		8.367
	Age		34.594
	Men		0.477
<i>IFLS5</i>	Worked Probability	31,748	0.662
	Education		9.043
	Age		35.617
	Men		0.480

Source: IFLS1, IFLS2, IFLS3, IFLS4, and IFLS5

Based on the highest level of education achieved, someone who is educated above high school has the highest probability of working, followed by elementary education level. The lowest probability is achieved when a person's highest education is junior high. This is possible because at the age of 15-18 years some individuals still take high school education. Elementary education level in addition to having a high probability of work also has the largest frequency of around 37% of all data. This indicates that education is still quite low, especially if the added respondents who did not finish elementary school to about 45,9%. This condition also shows that workers with elementary education are still quite high in Indonesia.

TABLE III. DESCRIPTIVE ANALYSIS OF WORK PROBABILITY 1993-2014

Independent Variable	Probability					
	No Worked		Worked		Total	
	freq	%	freq	%	Freq	%
Year						
1993	5,717	43.48	7,432	56.52	13,149	100
1997	9,096	48.90	9,505	51.10	18,601	100
2000	9,954	41.83	13,844	58.17	23,798	100
2007	11,187	40.95	16,135	59.05	27,322	100
2014	10,735	33.81	21,013	66.19	31,748	100
Education						
No school	4,627	45.78	5,481	54.22	10,108	100
PS	16,003	37.68	26,467	62.32	42,470	100
JHS	10,667	48.78	11,201	51.22	21,868	100
SHS	12,380	41.61	17,373	58.39	29,753	100
College	3,012	28.91	7,407	71.09	10,419	100
Cohort						
Young						
15-29	24,440	55.51	19,586	44.49	44,026	100
30-39	8,795	29.89	20,628	70.11	29,423	100
Old						
40-49	5,464	26.79	14,933	73.21	20,397	100
50-65	7,990	38.47	12,782	61.53	20,772	100
Gender						
Men	12,060	22.34	41,932	77.66	53,992	100
Female	34,629	57.12	25,997	42.88	60,626	100
Status						
Married	28,459	34.99	52,874	65.01	81,333	100
No Married	18,230	54.77	15,055	45.23	33,285	100

Source: IFLS1, IFLS2, IFLS3, IFLS4, and IFLS5

Based on gender, men who worked were around 77.66% while women were around 42.88% (Table 3). The number of women who work has experienced a significant increase compared to men. The average education and average income of workers has increased from year to year. Women's education has increased higher than men (Table 4), as well as for the average income (Table 5). This shows a reduction in the inequality of education and income between men and women.

TABLE IV. AVERAGE YEAR OF WORKERS EDUCATION IN INDONESIA

Year	Average Year of Education		
	Men	Female	Total
1993	6.16	4.88	5.72
1997	6.99	6.19	6.68
2000	7.76	6.78	7.38
2007	8.62	8.13	8.44
2014	9.28	9.06	9.19

Source: IFLS1, IFLS2, IFLS3, IFLS4, and IFLS5

TABLE V. AVERAGE REAL INCOME OF WORKERS IN INDONESIA

Year	Average Income (Rp)		
	Men	Female	Total
1993	4,525,068	2,781,745	3,929,261
1997	4,777,493	3,074,796	4,128,121
2000	4,554,587	2,947,824	3,939,690
2007	5,737,948	4,010,764	5,101,132
2014	7,799,850	6,941,581	7,448,773

Source: IFLS1, IFLS2, IFLS3, IFLS4, and IFLS5

The results of various estimates using the OLS, Heckit and IV methods are shown in Table 6. The return value of 15.2% is the return value in 1993 for the OLS method. The estimated return in 1997 declined 1.8% compared to 1993 to 13.4%. In 2000 decreased 3.2% compared to 1993 to 12%. In 2007 decreased by 2.9% compared to 1993 to 12.3%. In 2014 decreased 4.5% compared to 1993 to 10.7%.

The estimation of return value with Heckit method is 12.7% in 1993. There was a decrease of 1.79% in 1997 compared to 1993 to 10.91%. In 2000 there was a decrease in return of 3.26% compared to 1993 to 9.44%. In 2007 there was a decrease of 2.94% compared to 1993 to 9.76%. A decline of 4.55% occurred in 2014 compared to 1993 to be 8.15%.

The author used family background instruments that are father education and maternal education. Both instruments result in a statistical test of F greater than 10 for the first stage regression, so the instrument is feasible for use in method IV. The estimated value of educational coefficients and work experience for method IV has a greater value than the OLS method.

TABLE VI. RESULTS OF ESTIMATED RETURN ON EDUCATION

	OLS	Heckit	IV Father Educ	IV Mother Educ
School	0.152*** (0.0028)	0.127*** (0.0043)	0.212*** (0.0072)	0.220*** (0.0085)
School*1997	-0.0177*** (0.0035)	-0.0179*** (0.0038)	-0.0216** (0.0103)	-0.0127 (0.0125)
School*2000	-0.0323*** (0.0034)	-0.0326*** (0.0036)	-0.0341*** (0.0079)	-0.0448*** (0.0089)
School*2007	-0.0290*** (0.0034)	-0.0294*** (0.0035)	-0.0368*** (0.0077)	-0.0518*** (0.0086)
School*2014	-0.0452*** (0.0033)	-0.0455*** (0.0034)	-0.0195** (0.0087)	-0.0103 (0.0102)
Experience	0.0429*** (0.0012)	0.0294*** (0.0020)	0.0544*** (0.0017)	0.0556*** (0.0019)
Experience ²	0.0006*** (0.0000)	0.0006*** (0.0000)	0.0006*** (0.0000)	0.0007*** (0.0000)
1997	0.298*** (0.0290)	0.300*** (0.0287)	0.285*** (0.0602)	0.244*** (0.0719)
2000	0.264*** (0.0282)	0.265*** (0.0272)	0.258*** (0.0525)	0.331*** (0.0579)
2007	0.348*** (0.0289)	0.347*** (0.0277)	0.331*** (0.0544)	0.452*** (0.0597)
2014	0.776*** (0.0290)	0.773*** (0.0272)	0.328*** (0.0708)	0.211** (0.0853)
Constant	12.98*** (0.0311)	14.38*** (0.172)	12.34*** (0.0625)	12.37*** (0.0751)

Table 5. Cont.

Control Var	Yes	Yes	Yes	Yes
Selected				
School		0.0266*** (0.0010)		
Age		0.0238*** (0.0003)		
Men		0.954*** (0.0081)		
constant		-1.220*** (0.0164)		
Mills lamda		-1.036*** (0.124)		
Observation	68,060	114,618	49,260	43,182
Obs selected		68,060		
R-squared	0.274		0.228	0.237
Wald Chi2		17,603.92		
Prob>chi2		0.000		
Weak (F test)			1,211.75	648.33
Weak (P-val)			0.000	0.000

Source: IFLS1, IFLS2, IFLS3, IFLS4, and IFLS5

IV. DISCUSSION

The estimation results of the return on education value obtained by the author are quite large especially when compared to developed countries. But these results support the findings of Psacharopoulos. These are relative differences in human capital in developing countries compared to developed countries [23]. The level of return on education services for developing countries is higher than that of developed countries [23,24]. This result is also consistent with a summary of several studies on the value of educational services from Bils and Klenow, which most of the estimated returns for developing countries are higher than developed countries [25].

The Heckit method is used to overcome bias selection in OLS models. It appears that Heckit's estimation method is lower than OLS, which denotes the bias due to sample selection. The bias size of this sample selection is around 2.5%. These results differ from the findings of Purnastuti et al. who did not find enough evidence of bias due to sample selection. The educational variables used dummy sets of graduates of educational level [17]. This is in contrast to my study who use the year of education as a measure of educational variables.

The IV method in this paper produces a higher estimate of OLS. The difference in the estimated return on education value is around 6%. This larger value of IV estimates is consistent with most of the previous researchers [5]. The difference in return on education is still within the range of empirical study results the previous studies [5,7]. Card concluded that estimates with IV provide higher values above 30%. The estimate IV method is greater due to the trend of estimating the average effect in the low group. This term is known as LATE (Local Average Treatment Effect) [26].

V. CONCLUSION

The use of the Mincer model with OLS provides an estimated value of education services of 15.8% in 1993 and 10.85% in 2014. The estimation results indicate that there is bias due to selected sampling, since the sample selection is

only wage earners. The two stage method introduced by Heckman was used to overcome this. This method is known as the Heckit method. The result of Heckit's estimation of educational services is smaller than the OLS estimate of 14.2% in 1993 and 8.9% in 2014.

Both methods (OLS and Heckit) in estimating the value of returns for education services still allow the value of bias because the education variable is not truly endogenous and there are unobserved variables that do not enter the model. In this study the author uses instrumental variables (IV) to overcome the bias due to endogeneity in education variables. The estimation of the return value of education services using IV is higher than the OLS or Heckit method, which is 21.2% in 1993 and 19% in 2014.

The return on education services obtained is greater when using method IV in accordance with Card's statement. Card states that the estimation results of coefficients with method IV tend to give a greater value than OLS. The estimated IV value is greater because it estimates the average effect among low groups, while OLS estimates are the average effect estimates among all people (LATE) [26].

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