

Examining the Relationship between Clean/Fossil Energy Consumption and Sustainable Development

Sen Qiao, Wei Jun Tong, Liang Zhan and Hsinghung Chen*

School of Business, Macau University of Science and Technology, Macau, China

*Corresponding author

Abstract—Clean Energy including natural gas, hydropower, and nuclear energy, as well as fossil energy including petroleum and coal are the most important energy consumption in China. Although important, the research on the relationship between clean/fossil energy consumption and sustainable development has only been done limitedly. Some papers have used Granger causality and employed panel co-integration test and panel error correction model to study the related topic. Thus, this paper expects to examine the causal relationship of clean/fossil energy consumption with economic growth, environmental issues, and social problems.

Keywords—component; Clean and fossil energy consumption; sustainable development; social development; economic growth

I. INTRODUCTION

China has shown sustained economic growth at an average annual rate of 9.0% since its economic reforms beginning in the 1980s (Yang, 2010) [27]. Although the average income per capita is expected to grow more than triple from 2005 to 2020, many evidences show that the economic expansion has induced a substantial increase in energy consumption (Apergis and Payne, 2010; Wang et al., 2011; Dergiades et al., 2013; Zhang et al., 2013) [1,23,7,28]. In fact, China consumed 2,972 million tons of oil equivalent of primary energy including 83.5% of fossil energy and 16.5% of clean energy, accounting for 23% of the world's total in 2014, and the consumption in China overpassed that in the USA since 2010 (BP, 2015) [5]. Netherlands Environmental Assessment Agency announced that carbon dioxide emissions of China had surpassed those of the US in 2006, making China the largest carbon emitter (BP, 2008; Ma and He, 2008) [3,18]. CO₂ emissions in China kept increasing from 7,945, 8,660, to 9,208 million tons in 2011, 2012, and 2013, respectively (BP, 2014) [4]. As a result, half of cities in China cannot meet air quality standards, and one third of its land is affected by acid rain (EIA, 2012) [9]. The substantial economic growth has resulted in negative side effects including environmental degradation (such as climate change, forest degradation through acid rain) and social problems (such as educational shortfalls and increasing crime rates caused by migration). One instrument to alleviate the negative socio-economic and environmental impacts is to integrate economic, environmental and social policy into a national sustainable development strategy (Klemmer, 1994; Li and Oberheitmann, 2009; Bhattacharyya, 2012) [14,17,2]. Thus, sustainable development strategy is necessary for China.

II. LITERATURE REVIEW

There are more than 100 definitions for sustainable development (GDRC, 2005) [10]. From Brundtland report, development that meets the needs of the present without compromising the ability of future generations to meet their own needs is the most frequently used definition for sustainable development (GDRC, 2005) [10]. Generally speaking, sustainable development is assessed from three dimensions: economy, ecology and social dimensions (Klemmer, 1994) [14]. The level of sustainable development for alternative energy has been measured by certain indicators, such as energy efficiency, carbon dioxide emission intensities and social fairness outcomes (Li and Oberheitmann, 2009) [17]. Sustainable development strategy in China is embedded in the efforts to transform China into a modern country with the need for harmony among the environment, society, and economy (Bhattacharyya, 2012) [2]. Then, the purpose of this paper is to examine, from the perspective of clean and fossil energy, the quantitative context of the economic development and subsequent growth-related environmental and social problems.

The energy consumption and economic growth nexus have attracted a considerable amount of academic interests since the emergence of the oil shocks in 1970s (Wolde-Rufael and Menyah, 2011; Paul and Uddin, 2011) [26,21]. After the seminal work by Kraft and Kraft (1978), whose findings showed that it was difficult for all countries to carry out a unique energy policy recommendation for economic growth, the relevant literature has provided mixed and conflicting evidence with respect to this relationship [15]. For instance, neutrality hypothesis indicates that if there is no causality between these two variables, economic growth will be influenced by energy conservation policy with small effect or even no effect. However, the so-called growth hypothesis is established when uni-directional causality runs from consumption to growth. Under this circumstance, the link between energy consumption and economic growth is negative. On the other hand, the conservation hypothesis holds if causality runs only from GDP growth to energy consumption. Thus, the energy conservation policy may be adopted without considering the negative effect on GDP growth. Lastly, in the case of bi-directional causality, the feedback hypothesis suggests that it is significant to take the interdependent relationship between energy consumption and growth into consideration when drawing up policies. There is plenty of research in this regard, and the identification of the relationship at the disaggregated level becomes important because the effect of economic growth varies with different types of energy (Kum

et al., 2012) [16]. A new set of literature has appeared in order to study the relationship of various types of energy consumption, including electricity, coal, natural gas and gasoline, with economic growth (Omu et al., 2013) [20].

III. METHODOLOGY

A. Proposed Hypotheses

In order to examine the causal relationship of Clean/Fossil Energy consumption with economic growth, environmental issues and social problems, the following hypotheses are made, while the paper will employ panel data models to simulate the multi-variable regressions.

Many studies have shown that a substantial increase in energy consumption, especially in renewable energy consumption, has a significant relationship with economic development (Apergis and Payne, 2010) [1].

Hypothesis (a): The causal relationship of natural logarithm of real GDP per capita with change of clean energy consumption should be positive in a specific period.

Empirical evidence suggests that there is a bi-directional causal relationship between industrial and residential energy consumption to real GDP (Tsani, 2010) [22].

Hypothesis (b): The causal relationship of natural logarithm of real GDP per capita with change of fossil fuel-based energy consumption should be positive in a specific period.

Menegaki (2011) used annual panel data ranging from 1997 to 2007 for 27 European countries to examine the causal relationship of real GDP per capita with greenhouse emissions equivalents with base year 1990, and concluded that their relationship was apparently opposite [19].

Hypothesis (c): The causal relationship of natural logarithm of real GDP per capita with change of GHG emissions should be negative in a specific period.

The results indicate that a natural gas boom should increase salary income because of the combination of a greater demand for labor, an increase in the number of jobs, and a raise in the rent payments to private and public resource owners (Weber, 2012) [25].

Hypothesis (d): The causal relationship of natural logarithm of real GDP per capita with change of income per capita should be positive in a specific period.

The existence of supporting schemes for renewable energy represents an opportunity to reduce the risk by ensuring a fixed income for a certain period of time, and reflects the need to ensure the interests from private investors for technological innovation (Carneiro and Ferreira, 2012) [6].

Hypothesis (e): The causal relationship of natural logarithm of real GDP per capita with change of real gross fixed capital formation should be positive in a specific period.

Wolde-Rufael and Menyah (2011) examined the causal relationship of the change of GDP per capita with the change of

labor force and other parameters, and concluded that the relationship was positive with significance [26].

Hypothesis (f): The causal relationship of natural logarithm of real GDP per capita with change of labor force should be positive in a specific period.

B. Research Methodology

In this study, panel data is applied due to three major reasons. First, each province in China has quite different cultural backgrounds and specific characteristics. The use of panel data can effectively control the heterogeneity of individual province. Second, since only 15 provinces/places have sufficient data, there are not enough observations for regression. The employment of panel data can solve the problem. Third, from the perspective of econometrical study, panel data can provide us more degrees of freedom, more information, more variation, and less collinearity (Wang, 2008) [24]. Accordingly, the paper will employ panel data models to simulate the multi-variable regressions. From the proposed hypotheses, we have the following equation:

$$y_{i,t} = \alpha_i + \sum_i^k \beta_{i,t} X_{i,t} + \delta_{i,t}$$

where $y_{i,t}$ = natural logarithm of real GDP per capita in province i during t and $t-1$;

α_i = a regression constant for province i ;

$\beta_{i,t}$ = a regression coefficient in province i at t ;

$X_{i,t}$ = independent variables such as natural logarithm of fossil-based energy consumption or change of labor force; and

$\delta_{i,t}$ = a random residual in province i at t .

However, because the available time span is only 10 years, a simpler equation is employed to obtain the overall rather than individual relationships among variables:

$$y_{i,t} = \alpha_i + \sum_i^k \beta_i X_{i,t} + \varepsilon_{i,t}$$

Then, the proposed equation can be further represented by the following equation:

$$y_{i,t} = \alpha_i + \beta_1 X_{1(i,t)} + \beta_2 X_{2(i,t)} + \beta_3 X_{3(i,t)} + \beta_4 X_{4(i,t)} + \beta_5 X_{5(i,t)} + \beta_5 X_{5(i,t)} + \varepsilon_{i,t}$$

where $\alpha_i = \alpha + \alpha^*$

$y_{i,t}$ = natural logarithm of real GDP per capita in province i during t and $t-1$;

$X_{1(i,t)}$ = change of GHG emissions in province i at t ;

$X_{2(i,t)}$ = natural logarithm of clean energy consumption in province i at t ;

$X_{3(i,t)}$ = natural logarithm of fossil-based energy consumption in province i at t ;

$X_{4(i,t)}$ = change of income per capita in province i at t ;

$X_{5(i,t)}$ = change of fixed capital formation in province i at t ;

$X_{6(i,t)}$ = change of labor force in province i at t ;

$\varepsilon_{i,t}$ = a random residual in province i at t ;

α = average value of cross-section item in province i ; and

α^* = individual deviation from average value of cross-section item in province i .

To examine the presence of unit roots in a series, Levin-Lin-Chu (LLC) test, augmented Fisher Dickey-Fuller (Fisher-ADF) test and Fisher Phillips-Perron (Fisher-PP) test are well accepted methods. When co-integration test is applied to verify the stability of a system, Kao residual test is well accepted (Kao and Chiang, 2000) [13]. When employing the panel data model, three alternatives are pooled ordinary least square (OLS) regression model, fixed effect (FE) regression model, and random effect (RE) regression model. When likelihood test is passed, the FE regression model is more appropriate. When Hausman test is passed, the RE regression model is more suitable. Otherwise, the OLS regression model can be used. However, if both likelihood and Hausman tests are passed (in the paper, the number of provinces (N) is greater than the number of year (T), the FE model is more suitable if panel data is not randomly sampled (Hausman, 1978; Green, 2003) [12,11]. Oppositely, the RE model is more appropriate if panel data is randomly sampled.

Before simulating the regressions, several potential econometric problems need to be checked. The first one is collinearity between explanatory variables. People may have good reason to suspect a collinearity relation between variables, such as labor force and average income, because they may impact each other. This may be a critical problem if we employ just cross-sectional data. As mentioned above, panel data model can basically solve the collinearity problem. The second problem is heteroscedasticity among cross-sectional data. To solve this problem and possible cross-sectional correlation problem, we employ regressions that report Driscoll-Kraay standard errors. Such kinds of regressions are more robust if heteroscedasticity and cross-sectional correlation problems exist (Driscoll and Kraay, 1998) [8]. In addition, because the panel data in the case study has the characteristic of $N > T$, the problems of serial correlation and panel unit root may be reasonably ignored.

IV. DATA COLLECTION AND ANALYSIS

This study employs panel data regression model to examine the relationship of natural gas consumption with sustainable

development in China. Panel data sets include one dependant variable, real GDP per capita (unit: RMB), and six independent variables: clean energy consumption including hydropower, nuclear, and natural gas energy (unit: million tons of oil equivalent), fossil-based energy consumption including coal-based, and gasoline-based energy (unit: million tons of oil equivalent), GHG emissions (million tons), income per capita (unit: RMB), fixed capital formation (unit: billion of RMB) and labor force (unit: 10,000 persons). The data of 15 places in China are collected from 2005 to 2014. The sampled provinces and cities include XINJIANG, QINGHAI, GANSU, SHANXI, SICHUAN, HAINAN, GUANGDONG, HUBEI, HENAN, SHANDONG, JIANGSU, HEILONG, JILIN, LIAONING, and HEBEI. The paper selects the places having the most integrated data from 2005 through 2014. The descriptive statistics and covariance analysis of all variables are presented in Table I and II, respectively. Unit root test is passed after the first order difference of all variables ($P < 0.01$) and the result is shown in Table III. Kao Residual Co-integration Test is passed with a value of P equaling 1.25%, and the result is shown in Table IV.

TABLE I. DESCRIPTIVE STATISTICS FOR ALL VARIABLES

	GDP	Clean	Fossil	GHG	Income	Capital	Labor
Mean	8.961	23.999	25.108	0.027	0.113	0.207	-0.017
Max.	10.952	26.534	26.856	0.675	0.205	0.654	0.030
Min.	5.966	20.410	22.262	-0.202	0.000	-0.107	-0.082
Std. Dev.	1.133	1.358	1.096	0.117	0.045	0.119	0.018
Skewness	-0.490	-0.103	-0.846	2.222	-1.223	0.519	-0.508
Kurtosis	2.664	2.438	3.352	10.520	4.424	4.427	3.652
Obs.	150	150	150	150	150	150	150
Cross sections	15	15	15	15	15	15	15

TABLE II. THE RESULTS OF COVARIANCE ANALYSIS

	GDP	Clean	Fossil	GHG	Income	Capital	Labor
GDP	1.000						
Clean	0.091*	1.000					
Fossil	0.593***	0.132*	1.000				
GHG	0.388***	0.061*	-0.147**	1.000			
Income	0.280***	0.040	0.173***	0.085**	1.000		
Capital	0.007*	0.016*	0.135**	0.024*	0.543***	1.000	
Labor	-0.116	0.000*	-0.069*	0.029*	-0.385***	-0.181***	1.000

V. DISCUSSION

Using the multivariable instrument, the paper concludes that the development of clean/fossil energy consumption integrated with the consideration of the reduction of pollutant emissions, the improvement of income per capita, the creation of new jobs, and the attraction of capital investment does make a significant contribution to sustainable development in China. In other words, the importance of these findings lies on their implications for governments to adopt strategic policies in supporting economic growth, optimizing energy structure,

ensuring energy security, as well as in strengthening environmental protection and improving living standards.

TABLE III. THE RESULTS OF UNIT ROOT TEST

	Level			1st difference		
	LLC	Fisher-ADF	Fisher-PP	LLC	Fisher-ADF	Fisher-PP
GDP	-4.96***	25.13	77.60***	-11.21***	88.59***	127.92***
Clean	-23.90***	51.40***	52.11***			
Fossil	-3.77***	17.81	15.46	-10.70***	98.90***	126.79***
GHG	-7.81***	54.57***	54.06***			
Income	-15.32***	138.65**	243.57***			
Capital	-6.19***	96.95***	141.34***			
Labor	-8.72***	79.81***	104.18***			

TABLE IV. THE RESULTS OF COVARIANCE ANALYSIS

Kao Residual Co-integration Test		
Null Hypothesis: No co-integration		
	t-Statistic	Prob.
ADF	-2.241475	0.0125
Residual variance	0.017783	
HAC variance	0.030174	

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