

The Research of Passenger Flow Through Checkpoint Based on Queuing Theory

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Abstract. Long lines emerging recently at American airports incurred sharp criticism for the TSA, how to solve the problem becomes a concern both for the TSA and passengers. To explore the flow of passengers through a security checkpoint, propose some measures to modify the current process and some advice for managers, we develop modeling methods based on queuing theory. Queue length l , queuing probability c and average waiting time w can be obtained by solving equations. We then explore the relations between these variables and the passenger flow λ and give an expression $\lambda = w \cdot c / l$.

1. Introduction

Recently, lines in America seems to be insane, which frustrates many travelers. A recent survey conducted among American on the Internet says that compared with queuing at airports, 35% of participants prefer to do shopping, 22% prefer to clean up the mess after holiday picnics, and 18% would like to sleep in the sofa^[1]. So it is not difficult to see passengers' anxiety considering hours-long security wait time. This algorithm is also a part of our model regarding the passenger flow through the checkpoint. We set up parameters λ : passenger arrival rate, μ : service rate, and utilization rate of the service desk ρ as criteria.

2. Research results based on Queuing Theory:

Poisson process is a special random process which is frequently used to describe customer arrival disciplines in queuing theory. The probability of customers arriving in from t_1 to t_2 can be described by following formula, on the premise that the number of customers is n :

$$Pn(t_1, t_2) = P\{N(t_2) - N(t_1) = n\} \quad (t_1 < t_2; n \geq 0) \quad (1)$$

where

$N(t)$ is the number of customers arriving in from 0 to t

$Pn(t_1, t_2)$ follows the Poisson distribution with the parameter λ_1 when it meets one of the three following conditions: (1) stability (2) independence (3) universality^[2].

Through a series of deduction, the probability of n customers arriving in during the t interval $P_n(t)$ is described as:

$$P_n(t) = \frac{(\lambda_1 t)^n}{n!} e^{-\lambda_1 t} \quad (t > 0; n = 0, 1, 2, \dots) \quad (2)$$

so we can see $N(t)$ follows the Poisson distribution, its mathematical expectations and variances are as follows:

$$E(N(t)) = \sum_{n=0}^{\infty} n \frac{(\lambda_1 t)^n}{n!} e^{-\lambda_1 t} = \lambda_1 t, \quad \text{Var}(N(t)) = \lambda_1 t \quad (3)$$

especially, when $t = 1$, we can get $E(N(1)) = \lambda_1$, which represents average number of arrivals per unit time, also called the arrival rate^[3].

If the probability density function of random variable T is expressed as

$$f(t) = \begin{cases} \mu e^{-\mu t} & (t \geq 0) \\ 0 & (t < 0) \end{cases} \quad (\mu > 0) \tag{4}$$

then T follows the negative exponential distribution with parameter μ .

The distribution function is described as

$$F(t) = \begin{cases} 1 - \mu e^{-\mu t} & (t \geq 0) \\ 0 & (t < 0) \end{cases} \quad (\mu > 0) \tag{5}$$

next

$$E(T) = \frac{1}{\mu} \tag{6}$$

$$Var(T) = \frac{1}{\mu^2} \tag{7}$$

are easy to obtain. Where μ stands customers who have received service and leave the system per unit time, also called the average service rate of each check-in desk.

Assuming that each check-in desk work independently and average service rate is equal, namely $\mu_1 = \mu_2 = \dots = \mu_c = \mu$. So average service rate of the whole service agent is either $c\mu$ when $n \geq c$ or $n\mu$ when $n < c$. We define average utilization of service agent as

$$\rho = \frac{\lambda_1}{c\mu} \tag{8}$$

to prevent an infinite queue, ρ needs to satisfy $\rho < 1$. Fig.1 shows the process.

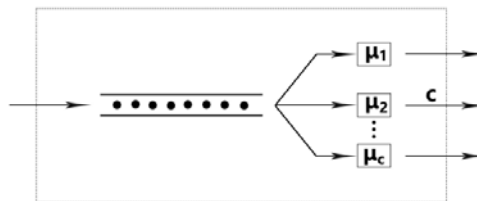


Fig.1 a sketch of service desks

Begin analysis of the queuing system from investigating the relation of system states. The transfer relations between these states are shown as figure 1.

The transition from state 0 to state 1 means that a customer has received service and leaved, the state transition probability is represented by μP_1 . The transition from state 2 to state 1 means that one of customers served at the service desk μ_1 and μ_2 , and the state transition probability is $2\mu P_2$.

Similarly, consider the case where the state transfer from state n to state $n - 1$. The state transition probability is $n\mu P_n$ in the case of $n \leq c$. While if $n > c$, limited number of service desks determine that there are c customers being at service at most, and $n - c$ customers waiting, the state transition probability is therefore changed into $c\mu P_n$.

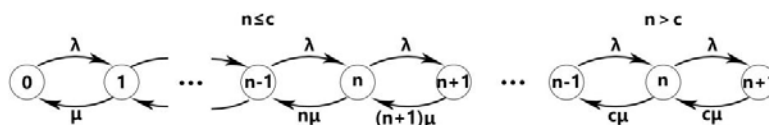


Fig.2 state transition process

From figure 2 we can conclude that

$$\begin{cases} \mu P_1 = \lambda_1 P_0 \\ (n+1)\mu P_{n+1} + \lambda_1 P_{n-1} = (\lambda_1 + n\mu)P_n & (1 \leq n \leq c) \\ c\mu P_{n+1} + \lambda_1 P_{n-1} = (\lambda_1 + c\mu)P_n & (n > c) \end{cases} \tag{9}$$

Here $\sum_{i=0}^{\infty} P_i = 1$ and $\rho \leq 1$

Solving the deterministic difference equations using recursion gives state probability as follows

$$P_0 = \left[\sum_{k=0}^{c-1} \frac{1}{k!} \left(\frac{\lambda_1}{\mu} \right)^k + \frac{1}{c!} \frac{1}{1-\rho} \left(\frac{\lambda_1}{\mu} \right)^c \right]^{-1} \quad (10)$$

$$P_n = \begin{cases} \frac{1}{n!} \left(\frac{\lambda_1}{\mu} \right)^n P_0 & (n \leq c) \\ \frac{1}{c! c^{n-c}} \left(\frac{\lambda_1}{\mu} \right)^n P_0 & (n > c) \end{cases} \quad (11)$$

system running indexes are derived as:

average queue length

$$\begin{cases} L_s = L_q + \frac{\lambda_1}{\mu} \\ L_q = \sum_{n=c+1}^{\infty} (n-c) P_n = \frac{(c\rho)^c \rho}{c!(1-\rho)^2} P_0 \end{cases} \quad (12)$$

average waiting time and sojourn time derived from the Little formula are shown as follows respectively

$$W_q = \frac{L_q}{\lambda_1}, \quad W_s = \frac{L_s}{\lambda_1} \quad (13)$$

The results are given in terms of requirements and reality, intermediate data processing are omitted for simplicity, the results are given as follows

average queue length

$$\begin{cases} L_q = \frac{\rho^{k+1}}{(k-\rho) \sum_{n=0}^{k-1} C_n \rho^n} \\ C_n = (k-1)! \frac{n-k}{n!} \end{cases} \quad (14)$$

average waiting time

$$W_q = \frac{L_q}{\lambda_1} \quad (15)$$

queuing probability

$$C(\rho, k) = \frac{k-\rho}{\rho} L_q \quad (16)$$

3. References

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